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### ***Research Materials***

## **An evaluation of the one percent clustered sample of the 1990 census of China**

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## **An evaluation of the one percent clustered sample of the 1990 census of China**

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### **Abstract**

We describe and evaluate a one percent clustered sample of the 1990 Census of China, the largest publicly-available micro sample of any Chinese census, using direct inspection as well as comparisons with published data drawn from the complete enumeration. In the absence of official documentation, we elucidate the basis of the clustering; detect duplicate cases; report corrected totals; and make comparisons between the sample data and tabulations based on the complete enumeration at the province and county levels. Although the sample contains several anomalies, we conclude that it is broadly serviceable.

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## 1. Introduction

Two micro-samples of the 1990 Chinese Census have circulated in China and abroad. The first, in order of creation, is a one percent sample of rural administrative villages (*xingzheng cun*) and urban neighborhoods known as street committees (*jiedao weiyuanhui*).<sup>3</sup> The second is a one percent sample of households. We refer to the former, the subject of this article, as the “one percent clustered sample,” and to the latter as the “one percent household sample.” These data sets are not public use micro samples (PUMS) in the sense understood by users of, for example, U.S. Census data. No official printed documentation has been released that is specific to either sample, and, in particular, there is no documentation or evaluation of the method and procedures used to draw the samples.

Recently the China Population Information and Research Center has made the one percent clustered sample available through the Texas A&M China Archive.<sup>4</sup> The one percent clustered sample is thus the only publicly-available micro sample of the 1990 census; indeed, it is largest readily available micro sample of any Chinese census.<sup>5</sup> We obtained the clustered sample and corresponding mortality data separately, from unofficial sources. Other researchers have obtained the clustered sample independently and have made use of it (e.g., Fan and Huang 1998; Li and Zhu 2000; Huang 2001). Our clustered sample is identical to the one purveyed by Texas A&M, except that the latter does not contain mortality data.

This paper evaluates the one percent clustered sample, using direct inspection as well as a series of comparisons with published data drawn from the complete enumeration. We first discuss the nature of the clustering, and report what we know about the sampling of clusters. We then note the existence of duplicate cases, and report corrected totals when duplicates are dropped. Subsequent sections describe geographic coverage of the one percent clustered sample, compare sample to published 100 percent tabulations for basic descriptors, and report selected comparisons between sample and 100 percent enumeration data at the county level.

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<sup>3</sup> An administrative village consists of one or several adjacent natural villages, and is the lowest level rural civil unit. In 1990, a neighborhood (street committee) was the corresponding urban civil unit.

<sup>4</sup> As of August 1, 2006, the data could be downloaded from The China Archive web site <http://chinaarchive.tamu.edu/portal/site/chinaarchive/>.

<sup>5</sup> A 1 per 1,000 sample of the 1982 Chinese census is available through IPUMS International <http://www.ipums.umn.edu/>. No micro sample of the 2000 census has been publicly released.

## 2. Clustering

Lacking any documentation on the method used in sampling, our description of the clustered sample relies on inference and a bit of hearsay. We have been told that the clustered sample was drawn as a way to provide a timely preview of census results, before final tabulations based on the complete enumeration could be prepared. We suspect, but cannot verify, that the sample is a systematic selection of all of the living persons in every hundredth administrative village (or urban neighborhood) within each province, and of all deaths in the sampled places that occurred in the 18 months leading up to the official July 1, 1990 date of the census.<sup>6</sup> Like the 1990 Census itself (except for national totals published in selected documents, e.g., State Council Population Census Office (1993)), the clustered sample contains civilians only.

Administrative villages and urban neighborhoods lend themselves as sampling units because they also serve as census districts, and census returns are sent up the line bundled by village or neighborhood. Sampling and data entry presumably took place in provincial offices. These circumstances may explain some anomalous aspects of the sample.

## 3. Duplicates

Duplicate cases are one such anomaly. Duplicates appeared in approximately half the provinces, and in all cases entire sample villages were duplicated. We considered the possibility that the duplicates were inserted intentionally, perhaps as a weighting scheme. But because the sample is far more faithful to the 100 percent tabulations (State Council Population Census Office 1993) when the duplicates are omitted, it appears more likely that the duplicates resulted from errors in data processing. In their discussion of the one percent clustered sample Li and Zhu (2000:228) also conclude that the duplicates are due to processing errors. We have removed the duplicates from all computations reported in this paper. Users should note that the Texas A&M data set also contains the duplicates.

Without duplicate records the sample consists of 8,514 administrative villages or urban neighborhoods containing 11,475,104 enumerated persons, which averages to 1,348 persons per sample unit. In addition there are 99,196 records of persons who died in the 18 months prior to the census.

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<sup>6</sup> According to the codebook, "this dataset (1% sampling) was prepared taking villages as sampling unit ..." (State Statistical Bureau 1994).

#### **4. Geographic coverage**

The clustered sample contains data in all 30 provinces and regions covered by the 1990 census. We consider first the extent to which the clustered sample reproduces the distribution of population across provinces and the municipalities with provincial status in 1990 (Beijing, Tianjin, and Shanghai). Table 1 shows the percentage distribution of population by province in the complete census, and the ratio of the sample to complete census percentage in each province. A ratio greater than one indicates that a province is oversampled. A ratio less than one indicates that a province is undersampled. The ratio varies from a low of .82 in Ningxia to a high of 1.29 in Tianjin Municipality, two provincial level units with small populations. Because the sample units are clusters, the sampling variability tends to be greater than would be expected for a simple random sample of individuals, and the extreme ratios occur in provinces with very small proportions of the population. The ratio in larger provinces generally varies between .95 and 1.05.

There is also broad geographic coverage within provinces. Approximately 91 percent of China's 2,845 county-level units contain at least one sample administrative village or urban neighborhood. Among the 2,600 county-level units with coverage, there is an average of 3.2 sample units per county. Sample coverage is shown in the accompanying map of China's counties (Figure 1). Counties containing at least one village unit in the sample are shown in gray or black, while counties with no coverage are in white. Coverage is quite regular in China Proper and the northeast, as contrasted with the sparser coverage in the Inner Asian regions of Inner Mongolia, Xinjiang, and Tibet, and parts of Gansu and Qinghai provinces. As may be seen in Table 1, none of these provinces is under-sampled. The sparseness of populations in these areas appears to explain the lack of coverage. Tibet, however, is an exception.

The sample for Tibet lacks cities and towns and is thus entirely rural. According to the official 100 percent tabulations for 1990 (Tibet Autonomous Region Population Census Office 1992), the Tibet Autonomous Region is 11.5 percent urban. The Tibet sample consists of 23 villages distributed over 10 of 78 possible counties. Thirteen villages are in a single county (Dingri, the site of Mount Everest), two are in one county, and the remaining eight are distributed one per county. It is likely that the sample villages not in Dingri are composites of county sub-samples. This is suggested by the large size of these units (approximately seven times as large as the Dingri sample villages), and by their sample code numbers, each of which is "1." The Tibet sample thus appears to have been constructed according to different principles from the rest of the one percent clustered sample. Nonetheless, the sample data accord well with the 100 percent tabulations for rural Tibet (Tibet Autonomous Region Population Census

**Table 1: Comparison of the Provincial Distribution of Population in the 1% Clustered Sample to that in 100% Census Tabulations, China, 1990**

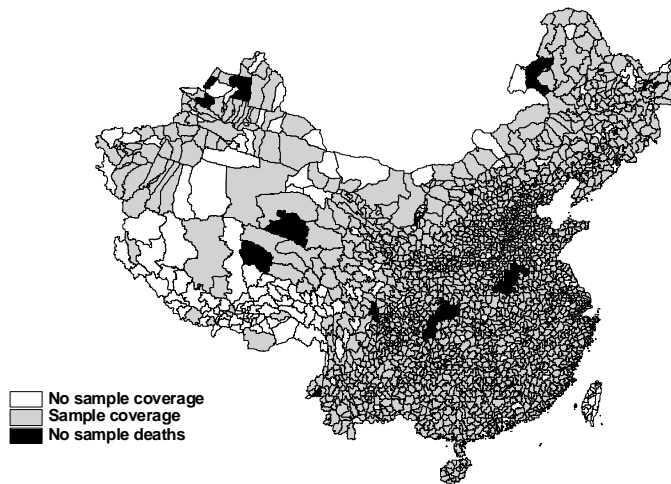
Province	Percent of Population in 100% Tabulations	Ratio of Sample to Census
Beijing	0.96	1.03
Tianjin	0.78	1.29
Hebei	5.40	0.96
Shanxi	2.54	1.03
Inner Mongolia	1.90	1.02
Liaoning	3.49	1.00
Jilin	2.18	1.07
Heilongjiang	3.12	1.00
Shanghai	1.18	1.14
Jiangsu	5.93	0.99
Zhejiang	3.67	0.99
Anhui	4.97	0.95
Fujian	2.66	1.11
Jiangxi	3.34	1.04
Shandong	7.46	0.97
Henan	7.57	0.96
Hubei	4.77	1.03
Hunan	5.37	1.02
Guangdong	5.56	0.98
Guangxi	3.74	1.03
Hainan	0.58	1.09
Sichuan	9.48	0.96
Guizhou	2.87	0.95
Yunnan	3.27	0.99
Tibet	0.19	1.09
Shaanxi	2.91	1.01
Gansu	1.98	1.06
Qinghai	0.39	1.27
Ningxia	0.41	0.82
Xinjiang	1.34	0.98
Total	100.01%	1.015

Sources: Population Census Office of China (1993) and one percent clustered sample.

Office 1992). For example, the 100 percent tabulation for Tibet shows that 76.3 percent of the rural population age 15 and above is illiterate. In the one percent sample, the corresponding figure is 77.4 percent. There is also a close correspondence with the rural age distribution, the distribution of rural women by their number of live births, and on other characteristics. The Tibetan sample may thus be useful for some purposes.

One other sample anomaly remains to be noted. The sample of dead persons includes four sample units that are in addition to the 8,514 sample units containing live persons. The four units, all in Guangzhou City, Guangdong Province, were apparently erroneously substituted for the death cases in four other Guangzhou sample units. These cases, 123 records in all, are included in the following comparisons.<sup>7</sup>

**Figure 1: County Coverage of One Percent Clustered Sample of the 1990 Chinese Census**



<sup>7</sup> The four dead-only units, all neighborhoods in the Liwan and Yuexiu Districts of Guangzhou Shi (GB 440103 and 440104 respectively), appear to be external to the original sample. This is suggested by the fact that both Liwan and Yuexiu are already fully represented by other neighborhoods containing live persons. Moreover, the Guangzhou Shi sample contains four neighborhoods with living persons but no corresponding dead persons, and three of these pertain to the same district (Liwan) that contains the dead-only units. This suggests an erroneous substitution of death records from non-sample neighborhoods. Thus, should the four dead-only units be excluded, it may be advisable to exclude the four living-only units of Guangzhou as well. The records from the dead-only units are included in the present assessment because their exclusion would have no appreciable affect on results at the national level. The 123 cases represent .0012 percent of the deaths in the total sample, and about 2.5 percent of the Guangdong Province sample. None of them pertain to persons age 0.



## **5. National comparisons of sample and census**

A series of comparisons at the national level (Tables 2-3 and Figures 2-3) reveals a reasonable concordance between the sample and the underlying complete census data as derived from published tabulations (State Council Population Census Office 1993). When the total of persons in the sample is multiplied by the reciprocal of the sampling fraction (i.e., 100) and divided by the census total, the resultant ratio is 1.02. The one percent clustered sample thus overstates the census population by two percent (see Table 2). Births are similarly overstated, while for the death sample (after alternative adjustments to account for excluded counties, discussed in the next section) the ratio of adjusted sample to the 100% is .992 to .998.

Having established a fair concordance between census and sample for total population, births, and deaths, we now consider the concordance of distributions of populations across various categories listed in Table 2. Most measures, such as percent male, percent rural, the sex ratio at birth, and deaths by semester, are within two percent of the census value. The distributions of population by occupation and by marital status are similarly close. There are two exceptions. The sample over-states the percent ever attending a university by 10 percent, perhaps as a consequence of the over-sample of the provincial level cities Shanghai and Tianjin that can be observed in Table 1. There is also an over-sampling of births in 1990 relative to 1989, for which we have no explanation.

Figure 2 presents a sample to census comparison of the sex-specific age distributions of those alive at the time of enumeration. The sample distribution of females by age is quite close to that for the census, varying within .5 percent at every age below 80. The male sample distribution is less regular. It contains an excess of males at ages 20-29, and a deficit of males age 60-75. Even so, these deviations are within one percent of the census value. The greater variability of males may be due to the greater concentration of males in sparse collective households and related institutional concentrations.

Figure 3, based on those who died in the 18-month period between January 1, 1989 and June 30, 1990, is constructed identically to Figure 2, but is based on age at death. The deviations of sample from census are more dramatic for the deceased. There is a notable dearth of dead males at ages 5-14—approximately 8 percent fewer than the corresponding census percentage. There is a similar dearth of females at ages 25-34. There are too many sample male death cases at ages 35-39, and too few sample female death cases at ages 45-49, approximately 10 percent fewer than expected. We have no explanation for these irregularities other than sampling variation.

**Table 2: Comparison of National Statistics Derived from the 1% Clustered Sample to Corresponding Statistics Derived from Complete Enumeration Tabulations of the 1990 Chinese Census**

Item	100% Census Value	Ratio of Sample to Census
Total population	1,130,510,638	1.02
Births (enumerated)	35,110,945	1.01
Deaths	10,328,899	.992 to .998 <sup>a</sup>
Population % male	51.47	1.00
Population % rural	73.80	0.99
Population % in collective households	2.89	1.02
Population % non-Han	8.08	1.00
Sex ratio of births (enumerated)	111.45	1.00
Sex ratio of births (mother reports)	114.18	1.00
Births by semester <sup>b</sup>		
Births 1989 first half %	32.25	0.98
Births 1989 second half %	37.88	0.98
Births 1990 first half %	29.87	1.04
Total	100.00	
Deaths total by semester <sup>b</sup>		
Deaths 1989 first half %	31.76	0.99
Deaths 1989 second half %	31.87	0.99
Deaths 1990 first half %	36.36	1.02
Total	99.99	
Educational Level <sup>b</sup>		
University %	0.62	1.10
Technical college %	0.97	1.04
Vocational high school %	1.74	0.97
Upper middle school %	7.30	1.01
Lower middle school %	26.50	1.00
Primary school %	42.27	1.00
Illiterate or semi-literate %	20.61	1.00
Total	100.01	

Table continued next page

**Table 2: Continued**

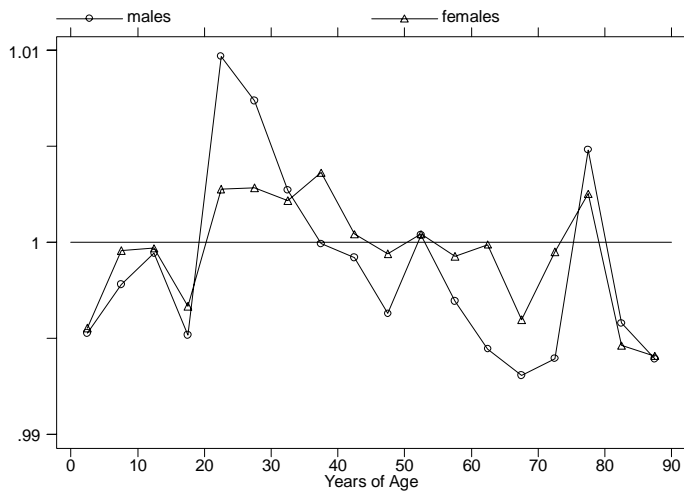
Item	100% Census Value	Ratio of Sample to Census
Occupation <sup>b</sup>		
Professional and Technical %	5.32	1.02
Cadres %	1.75	1.02
Administrative staff %	1.74	0.98
Commercial workers %	3.01	1.01
Service workers %	2.40	1.02
Agricultural workers %	70.61	0.99
Production workers %	15.17	1.02
Total	100.00	
Marital Status <sup>b</sup>		
Unmarried %	25.13	1.00
Married %	68.18	1.00
Widowed %	6.10	1.00
Divorced %	0.59	1.00
Total	100.00	

Sources: Population Census Office of China (1993) and one percent clustered sample.

<sup>a</sup> Sample deaths are adjusted for lack of death coverage in certain counties under different assumptions; see text and note 9 for an explanation.

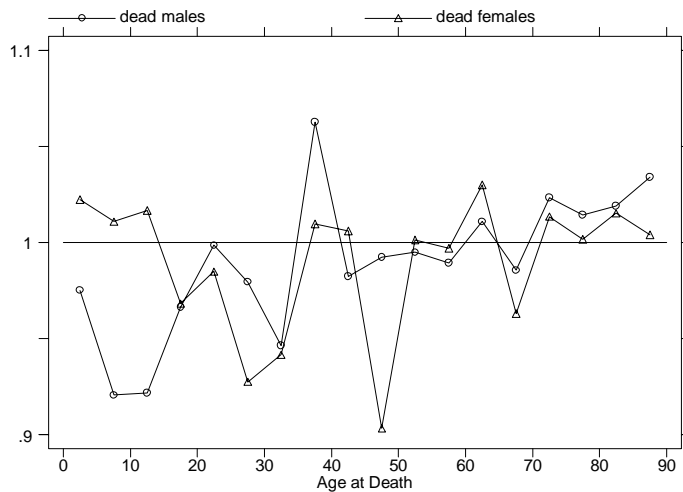
<sup>b</sup> Percentage distribution sums to 100 percent, with deviations due to rounding.

**Figure 2: Ratio of Sample to Census, Live Population, China, 1990.**



Sources: Population Census Office of China (1993) and one percent clustered sample.

**Figure 3: Ratio of Sample to Census, Dead Population, China, 1990**



Sources: Population Census Office of China (1993) and one percent clustered sample.

Because infant mortality is of particular interest, in Table 3 we further compare deaths at age 0 conditional on sex and semester of birth with the corresponding figures from published census tabulations. Official sources do not document the calculation of infant mortality *rates* with detail sufficient to sustain independent replication. For this reason our analysis is restricted to comparison of sample *frequencies* of infant death with those derived from the 100 percent enumeration tables (State Council Population Census Office 1993). The interior cells of Table 3 display ratios of sample infant deaths to complete enumeration infant deaths conditional on sex and semester of birth. The row and column margins contain ratios separately for sex and semester. Because of gaps in death coverage at all ages at the county level, which are discussed in the next section, the ratio of sample infant deaths to complete census infant deaths is less than one (.960), which suggests that the clustered sample undercounts infant deaths. However, upon adjusting for county coverage gaps in death reporting under the assumption that infant deaths were missed with probability identical to that for deaths to older individuals, the ratio of sample infant deaths to complete census infant deaths is found to be in the range of .991 to .999. This result suggests that infant deaths are not specifically undersampled in the one percent clustered sample. The sex-semester specific ratios in Table 3 should thus be considered to be downwardly biased owing to the absence of adjustment for nonreporting of deaths in particular counties. Of greater concern is the apparent undersampling of male relative to female infant deaths in every semester, as well as the inconsistency over semesters in the relative undersampling.<sup>8</sup> We have no explanation for this variability, but note that at a minimum it complicates the conclusions that can be drawn from individual level analyses of infant mortality.

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<sup>8</sup> We employed Poisson and negative binomial models, and checked for over-dispersion, to reach these conclusions. The sex effect is significant at the .1 level. Of the semester contrasts, only that between the first and third semester is significant at the .05 level.

**Table 3: One Percent Cluster Sample Deaths at Age 0 (Multiplied by 100)  
Divided by Corresponding 100% Enumeration Deaths,  
Conditional on Sex and Semester of Death, China, 1990<sup>a</sup>**

Semester	Male	Female	Total
1989: January-June	0.911 (120,406)	0.946 (118,893)	0.928 (239,299)
1989: July-December	0.959 (142,015)	0.962 (145,593)	0.961 (287,608)
1990: January-June	0.943 (177,418)	1.02 (185,177)	0.981 (362,595)
Total	0.941 (439,839)	0.980 (449,663)	0.960 (889,502)

Sources: Population Census Office of China (1993) and one percent clustered sample.

Note: Numbers in parentheses are death counts from 100 percent enumeration

<sup>a</sup>Sample values are unadjusted for lack of death coverage in certain counties; see text for discussion.

## 6. County-level comparisons of sample and census

For total population at the county level, if the sample data are unbiased, a regression of sample population counts on total enumeration population counts should return a coefficient that is approximately .01. For this calculation we found a slope of .0095 ( $N=2,600$ ). Fitting a cubic polynomial spline to the data revealed modest departures from linearity. If a handful of counties is excluded, the regression coefficient becomes .01. Thus, with respect to total population, the clustered one percent sample appears to be nearly unbiased at the county level.

We also compared total sample deaths to total census deaths at the county level. Here we detected a problem that required consideration prior to computation of a regression comparable to that for total population. Of the 2,600 county-level units in the clustered sample, 97 contain no death data. For a subset of these 97 counties, the absence of deaths is probably due to procedures followed at the local level, rather than to the inherent variability occasioned by probability sampling. In three contiguous prefectures in Henan (Shangqiu, Zhoukou, and Zhumadian), and two contiguous prefectures in Sichuan (Wanxian and Fuling), there are no mortality data. These five prefectures, containing 34,995,661 persons in 1990, account for 43 of the 97 county-

level units for which there are no deaths. The county-level sample sizes in these prefectures range from a minimum of 1,600 to a maximum of 15,441. Given the contiguity of the counties, and their correspondence to the areas of prefectural administrations, it is probable that the lack of mortality data in the 43 counties is due to some aspect of procedure and not sampling variability. In our judgment, these five prefectures should be excluded from any analysis of mortality.

The remaining 54 zero-death county-level units are geographically scattered, although many pertain to urban units in Heilongjiang and Anhui. They are far less populous on average than the 43 described above, together containing only 8,166,649 persons. It is at least theoretically possible that some or all of the sampled villages in these 54 county-level units recorded no deaths in the 18 months prior to the census. We addressed this question statistically, using total population as well as total death counts available in 1990 county census volumes. The structure of the inferential problem is to determine whether zero sampled deaths in each of the 54 counties is plausible or implausible, given a known death count and total population, and a sample population total, for each county. In our explorations, we concluded that zero sample deaths were implausible in all but nine of the counties. Other analysts might well use different methods and reach slightly different conclusions. For this reason, we present lower and upper bounds on the number of excluded counties. Specifically, at a minimum the 43 zero-death counties in five prefectures previously discussed should be excluded for the purposes of mortality analysis. At a maximum 97 zero-death counties should be excluded.

When zero-death counties are excluded, the complete-census death counts need to be reduced correspondingly, in order to make valid comparisons of sample-estimated deaths to complete-census deaths at the national level. The ratio of sample-to-census deaths in Table 2 provides this upper and lower bound for comparison. The lower bound is based on excluding 43 zero-death counties; the upper bound is based on excluding 97 zero-death counties.<sup>9</sup> The 97 zero-death counties are portrayed in the map (Figure 1) with black shading; in the complete census, these counties contributed 3.79 percent of deaths. The 97 zero-death counties are listed in Appendix 1.<sup>10</sup>

If death sampling is unbiased at the county level, then a regression of sampled deaths at the county level on complete census deaths at the county level should have a slope that is approximately 0.01. Omitting the 43 counties in five zero-death prefectures, we find a slope of .0096 ( $N=2,276$ ); omitting all 97 zero-death counties, we

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<sup>9</sup> The first sample to census ratio of deaths in Table 2 excludes only deaths from the 43 zero-death counties of the five prefectures, while the latter excludes deaths from all 97 zero-death counties. Counties in the five prefectures account for 3.177% of census deaths, while all 97 counties account for 3.794% of census deaths.

<sup>10</sup> The 43 counties in the five zero-death prefectures in Henan and Sichuan are distinguished in Appendix 1 by GB code (4123xx, 4127xx, 4128xx, and 5122xx, 5123xx respectively).

find a slope of .0095 ( $N=2,222$ ).<sup>11</sup> Thus, it appears, apart from the zero-death problem at the county level, that the one percent clustered sample is approximately unbiased for total deaths at the county level.

## **7. Discussion**

The one percent clustered sample appears to be a true one percent sample of the 1990 census. It reproduces the geographic distribution of population and major population components quite well. Although the clustering of the sample reduces precision, it permits contextual analyses based on multilevel methods of statistical analysis.

There are anomalies. The sample for Tibet lacks urban units and appears to use a different sampling procedure. The national distribution of deaths by age for males is irregular, and male infant mortality is somewhat under-sampled relative to the census. These deficiencies must be assessed for their relevance to specific analytic purposes. For example, there is mounting evidence (e.g., Ministry of Health 1999) that the 1990 census underreported infant mortality by a margin far wider than the gap between sample and census infant mortality. Against this kind of uncertainty, the sample can be useful. The results of the sample/enumeration comparisons we have presented suggest that the one percent clustered sample will be serviceable for many purposes.

## **8. Acknowledgements**

We gratefully acknowledge the valuable advice of Griffith Feeney and G. William Skinner, the research assistance of Yingying Zhou, and the support of the Henry Luce Foundation.

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<sup>11</sup> The base  $N$  for this regression is limited to 2,319 county-level units because we lack 100% death data on 281 units, mainly urban districts.



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## Appendix 1:

### Zero Death Counties Selected for Exclusion from Mortality Analysis in the One Percent Clustered Sample, China, 1990

GB Code	Province	Name
150103	Inner Mongolia	Huhehaote: Huimin qu
150122	Inner Mongolia	Tuoketuo xian
150402	Inner Mongolia	Chifeng: Hongshan qu
210802	Liaoning	Yingkou: Zhanqian qu
210803	Liaoning	Yingkou: Xishi qu
210811	Liaoning	Yingkou: Laobian qu
230402	Heilongjiang	Hegang: Xiangyang qu
230403	Heilongjiang	Hegang: Gongnong qu
230702	Heilongjiang	Yichun: Yichun qu
230705	Heilongjiang	Yichun: Xilin qu
230811	Heilongjiang	Jiamusi: Jiaoqu
230826	Heilongjiang	Huachuan xian
230834	Heilongjiang	Youyi xian
230881	Heilongjiang	Tongjiang shi
232603	Heilongjiang	Wudalianchi shi
320703	Jiangsu	Lianyungang: Lianyun qu
320704	Jiangsu	Lianyungang: Yuntai qu
320705	Jiangsu	Lianyungang: Xinpū qu
330921	Zhejiang	Daishan xian
340302	Anhui	Bengbu: Dong qu
340304	Anhui	Bengbu: Xi qu
340404	Anhui	Huainan: Xiejiaji qu
340503	Anhui	Ma'anshan: Huashan qu
340702	Anhui	Tongling: Tonggongshan qu
340803	Anhui	Anqing: Daguan qu
341002	Anhui	Huangshan shi CC: Tunxi qu
341003	Anhui	Huangshan: Huangshan qu

*continued next page*

**Appendix 1: Continued**

GB Code	Province	Name
341004	Anhui	Huangshan: Huizhou qu
341023	Anhui	Yi xian
342101	Anhui	Fuyang shi
342530	Anhui	Jingde xian
350203	Fujian	Xiamen: Siming qu
360302	Jiangxi	Pingxiang: Chengguan qu
360311	Jiangxi	Pingxiang: Shangli qu
362124	Jiangxi	Dayu xian
362129	Jiangxi	Dingnan xian
410411	Henan	Pingdingshan: Jiaoqu
412321	Henan	Yucheng xian
412322	Henan	Shangqiu xian
412323	Henan	Minquan xian
412324	Henan	Ningling xian
412325	Henan	Sui xian
412326	Henan	Xiayi xian
412327	Henan	Zhecheng xian
412328	Henan	Yongcheng xian
412701	Henan	Zhoukou shi
412721	Henan	Fugou xian
412722	Henan	Xihua xian
412723	Henan	Shangshui xian
412724	Henan	Taikang xian
412725	Henan	Luyi xian
412726	Henan	Dancheng xian
412727	Henan	Huaiyang xian
412728	Henan	Shenqiu xian
412729	Henan	Xiangcheng xian
412801	Henan	Zhumadian shi
412821	Henan	Queshan xian

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## Appendix 1: Continued

GB Code	Province	Name
412822	Henan	Biyang xian
412823	Henan	Suiping xian
412824	Henan	Xiping xian
412825	Henan	Shangcai xian
412826	Henan	Ru`nan xian
412827	Henan	Pingyu xian
412828	Henan	Xincai xian
412829	Henan	Zhengyang xian
450502	Guangxi	Beihai: Haicheng qu
512201	Sichuan	Wanxian shi
512221	Sichuan	Wan xian
512222	Sichuan	Kai xian
512223	Sichuan	Zhong xian
512224	Sichuan	Liangping xian
512225	Sichuan	Yunyang xian
512226	Sichuan	Fengjie xian
512227	Sichuan	Wushan xian
512228	Sichuan	Wuxi xian
512229	Sichuan	Chengkou xian
512301	Sichuan	Fuling shi
512322	Sichuan	Dianjiang xian
512323	Sichuan	Nanchuan xian
512324	Sichuan	Fengdu xian
512326	Sichuan	Wulong xian
513227	Sichuan	Xiaojin xian
533121	Yunnan	Luxi xian
610303	Shaanxi	Baoji: Jintai qu
620105	Gansu	Lanzhou: Anning qu
640121	Ningxia	Yongning xian
654225	Xinjiang	Yumin xian
654226	Xinjiang	Hebukesai`er Mengguzu zizhixian